

# New Bayesian approach for uncertainty computation at finite distance in nonlinear mixed effects models: simulation study and application

M. GUHL,<sup>①</sup> J. BERTRAND,<sup>1</sup> L. FAYETTE,<sup>1</sup> J. GUEJ,<sup>1</sup> F. MENTRÉ,<sup>1,2</sup> E. COMETS<sup>1,3</sup>

<sup>①</sup>melanie.guhl@inserm.fr, <sup>1</sup>Université Paris Cité, Inserm, IAME, F-75018 Paris, France, <sup>2</sup>AP-HP Hôpital Bichat, Département d'Epidémiologie Biostatistiques et Recherche Clinique, Paris, France, <sup>3</sup>Univ Rennes, Inserm, EHESP, Irset - UMR\_S 1085, F-35000 Rennes, France

## Introduction

**Non linear mixed effects models (NLMEM):** powerful tool to model longitudinal data

$\theta = c(\mu, \beta, \Omega, a, b)$ : vector of population parameters to estimate ( $\mu$ : fixed effects,  $\beta$ : covariate effects,  $\Omega$ : variance matrix of random effects,  $a, b$ : error parameters)

## Frequentist paradigm:

- Uncertainty of the maximum likelihood estimator (MLE) can be computed **asymptotically** through the standard errors (SE) based on the Fisher Information Matrix (FIM), and confidence intervals are built under a gaussian hypothesis (**Asympt**)
- When working at **finite distance**, **Asympt** underestimates the SE of NLMEM [1,2]
- Bootstrap or Sampling Importance Resampling (**SIR**) [3] methods have been proposed
- Sampling methods computationally heavy

Ueckert et. al. [4] proposed to compute uncertainty in a **Bayesian paradigm following the frequentist inference:**

- Under some regularity conditions on the prior, the limit distributions of the MLE and the maximum a posteriori (MAP) estimator are equivalent (Bernstein-von Mises theorem)
  - Posterior distributions are used to compute uncertainty, e.g. confidence intervals using quantiles
- Method implemented via Hamiltonian Monte Carlo algorithm in Stan (**Post**) [5,6]

**Objective:** To borrow from the **Bayesian paradigm** to estimate uncertainty through posterior distributions in NLMEM at finite distance, **in parallel of the frequentist estimation** in SAEM

## Methods

In the **saemix** R package [7], integration of a Bayesian step based on the Metropolis Hastings (MH) algorithm (**SAEM\_MH**) [8]

At each iteration  $k$  of SAEM at convergence phase, after simulation step/stochastic approximation/maximisation step

### Bayesian step:

Set a prior distribution  $p(\cdot)$  on  $\theta$

$\theta_{0_k}^{MH} = \theta_k$  the current frequentist estimation of SAEM

For  $m_k = 1, \dots, M_k$ :

- Draw  $\theta_{m_k}$  in a kernel distribution  $q(\cdot)$
- $\theta_{m_k}$  is accepted with probability  $\alpha$ :

$$\alpha = \min \left( 1, \frac{ll(\theta_{m_k}, \bar{\psi}_{m_k})p(\theta_{m_k})q(\theta_{m_k-1}^{MH})}{ll(\theta_{m_k-1}^{MH}, \bar{\psi}_{m_k-1}^{MH})p(\theta_{m_k-1}^{MH})q(\theta_{m_k})} \right)$$

with  $\bar{\psi}_{m_k}$  the mean of  $z = 50$  samples drawn from the conditional distribution  $\mathbb{P}(\psi_{m_k} | y, \theta_{m_k})$

Keep the last iteration  $\theta_{M_k}^{MH}$  of each MH chain to get a chain of  $K_2$  samples

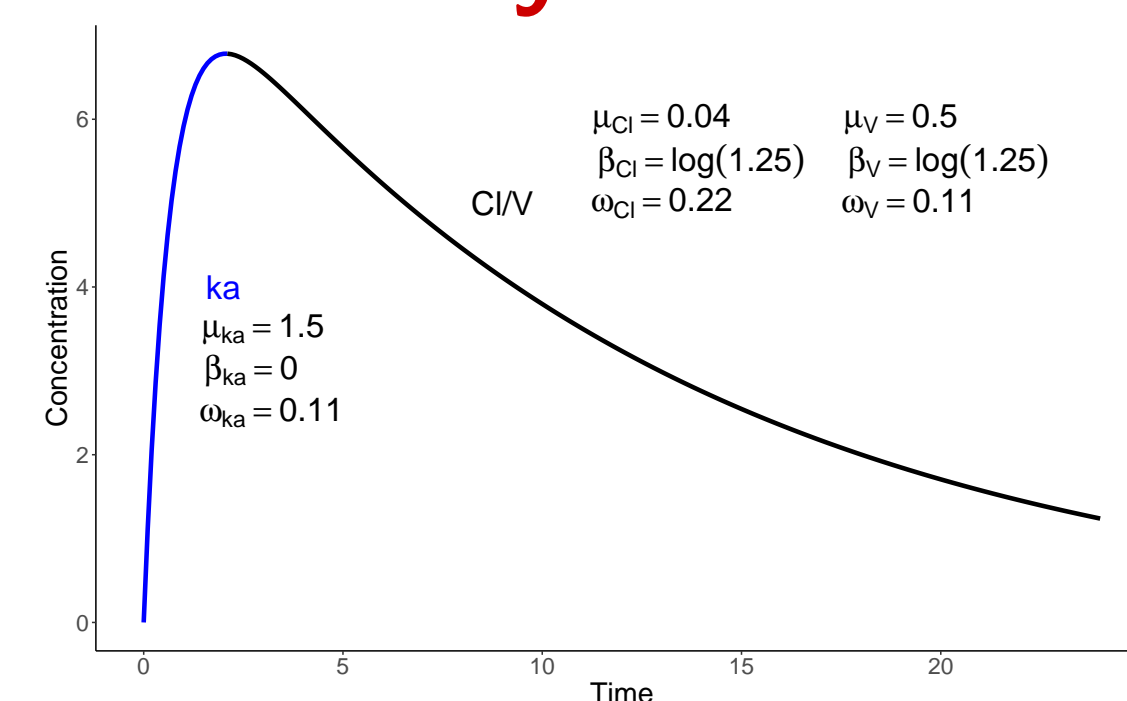
Parameters to be calibrated:

- $M_k$ , here  $M_k = 100$
- Prior distribution, here  $p = \mathcal{N}(\theta_p, sd_p)$  with  $\theta_p$  the simulated values of  $\theta$  and a coefficient of variation  $sd_p/\theta_p=30\%$  for  $\mu$  and  $50\%$  for  $\beta, \Omega, a$  and  $b$
- Kernel distribution, here  $q = \mathcal{N}(\theta_q, sd_q)$ ,  $\theta_q = \theta_k$  and  $sd_q = \text{inf}_q \times FIM_k^{-1}$ , with  $\text{inf}_q = 1, 1.5, 2$

## PK simulation study

- Inspired by theophylline data
- Proportional error model ( $b$ )
- Treatment groups of equal size

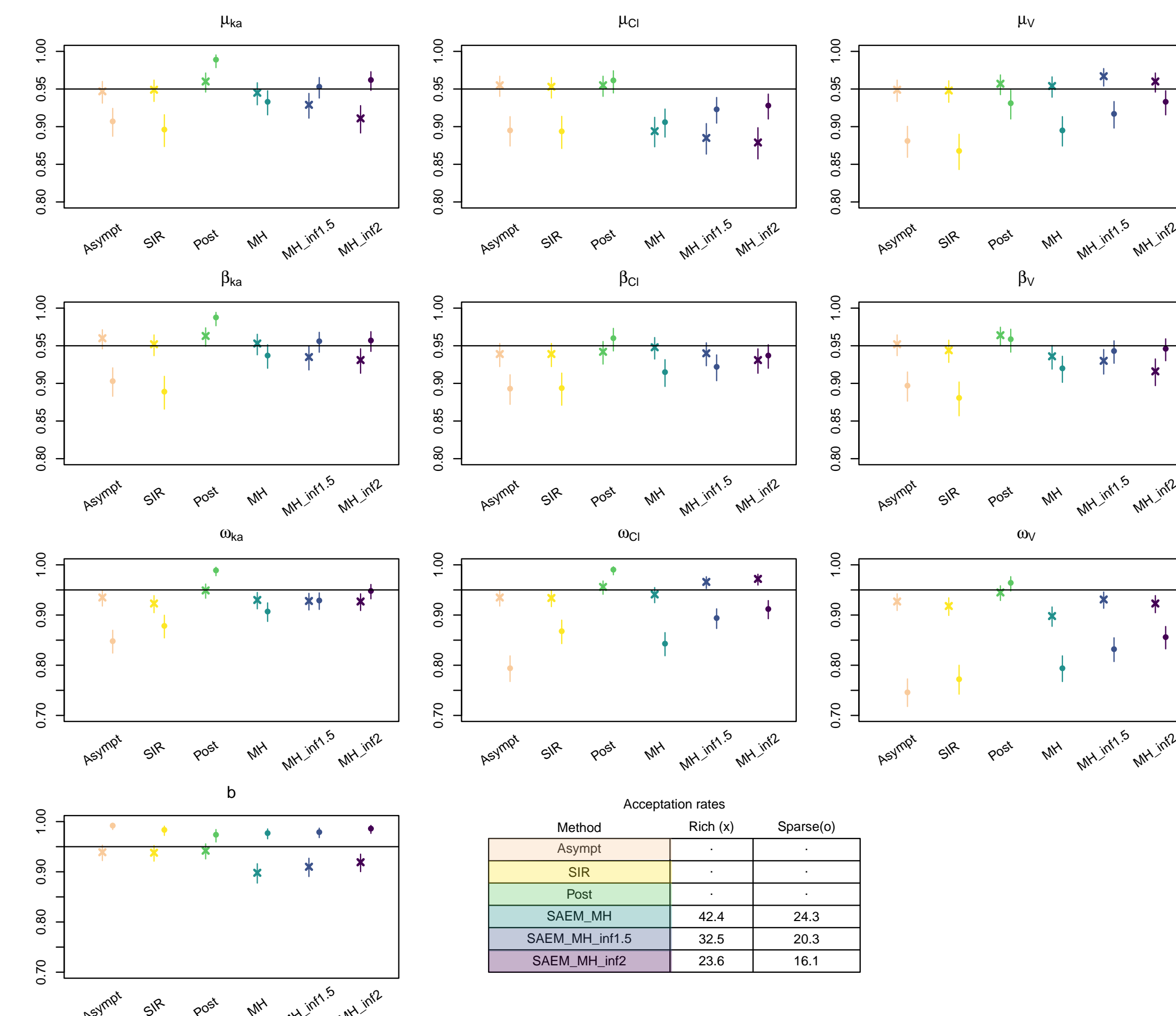
$$C(t) = \frac{\text{dose}}{V} \frac{k_a}{k_a - \frac{Cl}{V}} e^{(\frac{Cl}{V}t - k_a t)}$$



Two designs of 1000 simulated datasets:  
 Rich N=150 patients n=10 points  
 Sparse N=12 n=3

Evaluation:

- 10 population parameters ( $\beta_{ka}$  estimated)
- 95% coverage rates (CR) and their confidence intervals
- Acceptation rates (AR)
- Comparison with **Asympt**, **SIR** and **Post**



On N=150 and n=10 (x):

- All methods provide CR within the target range

On N=12 and n=3 (o):

- With **Asympt**, **SIR** and **SAEM\_MH**, CR systematically below the target
- Post** gives CR below or above the target
- SAEM\_MH\_2**, with inflated kernel variance by a factor 2, gives more controlled CR

In both cases, AR of MH between 15% and 40%: seems reasonable given the number of dimensions

Problem: when simulating additional data with increased variability and/or strong correlations, AR dropped

**SAEM\_MH** seems promising but needs further investigation on more challenging settings to higher the AR

Alternatives to the current MH algorithm currently investigated:

- Use AR as a guide for adaptive  $\text{inf}_q$  on  $sd_q$
- Adaptive kernel variance from literature [9]
- Block sampling of fixed effects and variance parameters separately
- Conditional univariate sampling (Gibbs sampling) with or without variance inflation for each parameter separately
- Random walks

## NEWS-2 modelling

- Discovery trial [10] (promoted by Inserm, PI: Pr Florence Ader): European clinical trial aiming to evaluate antiviral drugs for the treatment of Covid19
- Patients hospitalised for Covid19 followed for 29 days (patients recovering left the hospital but were followed up at day 15 and day 29)

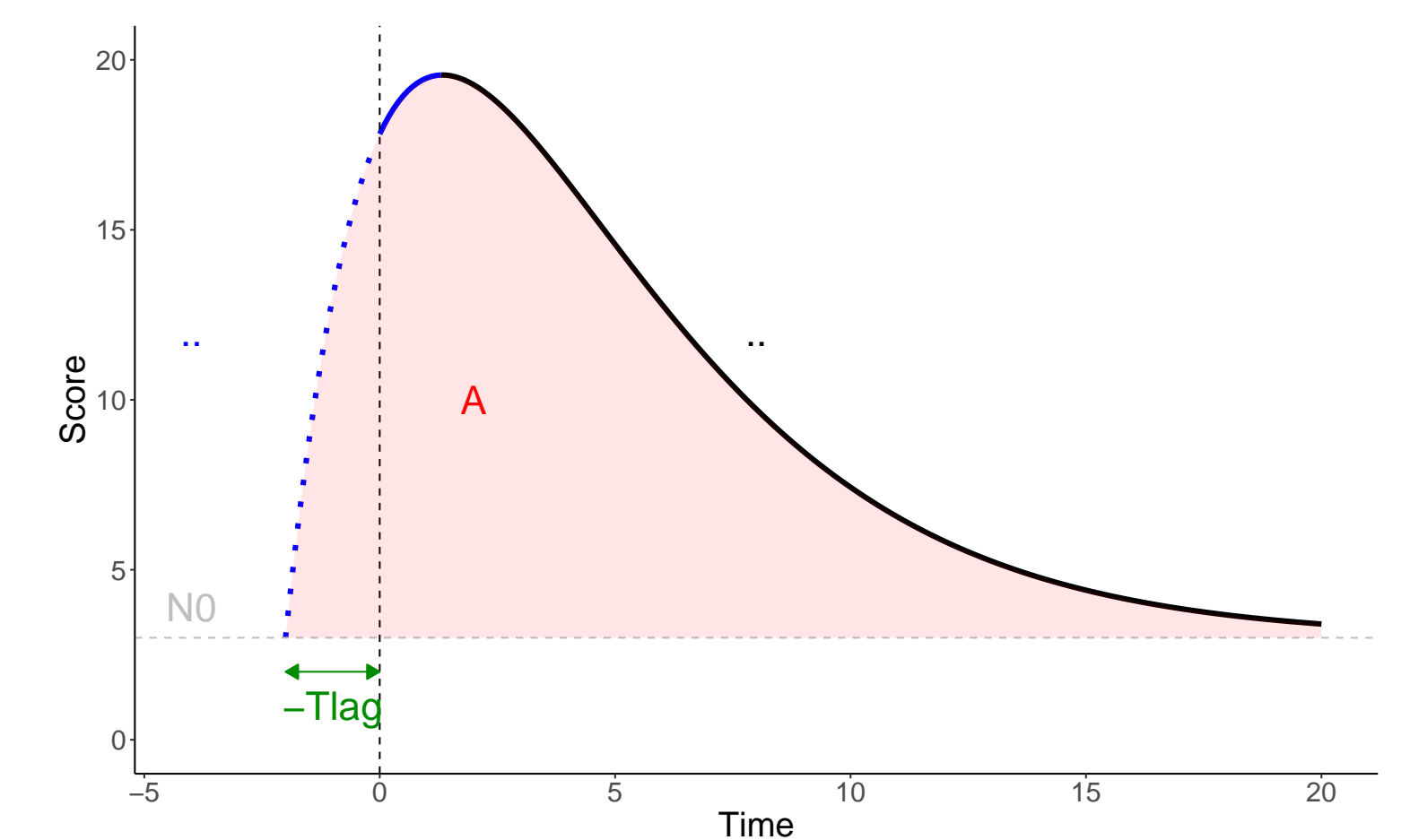
- NEWS-2 [11]: composite score ranging from 0 (best) to 20 (worst), used in emergency medicine to assess the clinical status of patients

Our objective: model the evolution of NEWS-2 and test for a treatment effect between:

- the standard of care (SoC, N=408 patients) arm
- the SoC + remdesivir (N=402 patients) arm

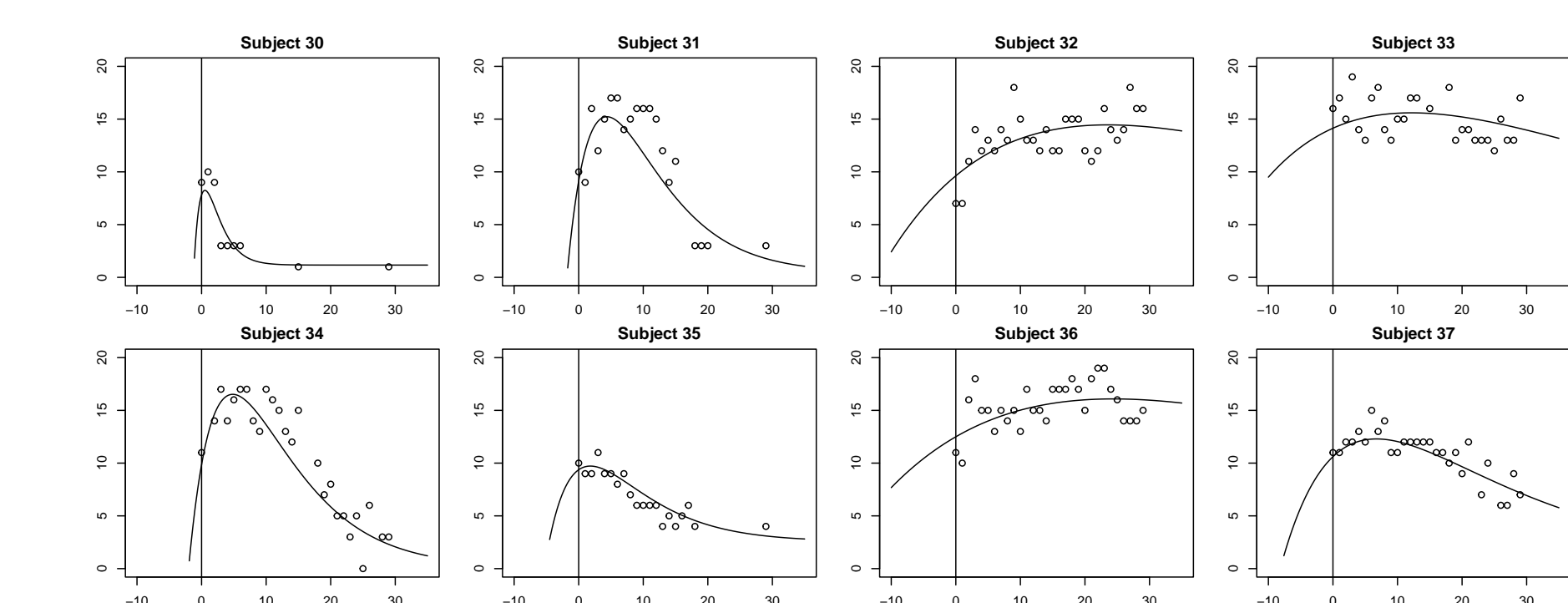
**Bateman function** to model the typical evolution of the score considered as continuous:

$$\text{NEWS2}(t) = N_0 + A \lambda^2 (t + T_{lag}) e^{-\lambda(t + T_{lag})}$$



- $T_{lag}$ ,  $\lambda$  and  $A$  follow a log-normal distribution
- $N_0$  follows a logit-normal distribution between 0 and 3
- log-transformation of the data and model as  $\log(1+\text{NEWS2})$  to account for the score being bounded by 0, with a constant error ( $a$ ) model

Parameter	Estimate	RSE (%)			
		Asympt	SIR	Post	MH
$\mu_{T_{lag}}$ (days)	2.85	10	9	7	13
$\beta_{age}(T_{lag})$	0.50	23	23	24	13
$\beta_{remde}(T_{lag})$	0.13	81	79	71	21
$\mu_A$ (days)	106.60	9	7	1	7
$\beta_{age}(A)$	0.56	18	18	16	6
$\beta_{remde}(A)$	-0.02	566	441	2084	12
$\mu_{N_0}$	0.18	19	19	16	29
$\beta_{cardiac}(N_0)$	0.21	174	175	196	57
$\beta_{diabetes}(N_0)$	1.01	38	40	52	42
$\mu_\lambda$ (days <sup>-1</sup> )	0.20	7	6	2	6
$\beta_{age}(\lambda)$	-0.53	16	16	15	5
$\beta_{remde}(\lambda)$	-0.03	245	151	233	12
$\omega_{T_{lag}}$	1.11	5	5	6	4
$\omega_A$	1.28	3	3	3	3
$\omega_{N_0}$	2.86	6	6	8	5
$\omega_\lambda$	1.05	3	3	4	3
$r_{T_{lag},\lambda}$	0.83	10	10	2	14
$r_{T_{lag},A}$	-0.90	10	11	3	15
$r_{A,\lambda}$	-0.97	8	8	0	9
$a$	0.25	1	1	1	2



→ With **SAEM\_MH** method:

- AR were critically low due to the challenging setting
- Calibration is needed to use **SAEM\_MH** on this data

## Conclusion

- As expected, **Asympt** is degraded at finite distance
- SIR** does not improve the results
- Post** gives CR above the target on sparse design and needs further work to investigate suitable priors
- SAEM\_MH** also needs further investigation for the calibration of the prior and kernel distributions on challenging settings

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